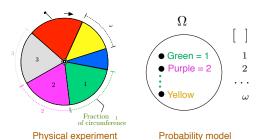
Lecture 16: Continuing Probability.

Wrap up: Probability Formalism.

Events, Conditional Probability, Independence, Bayes' Rule

Probability Space: Formalism

Physical model of a general non-uniform probability space:

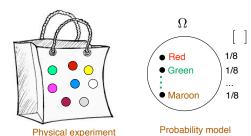


The roulette wheel stops in sector ω with probability p_{ω} .

$$\Omega = \{1, 2, 3, ..., N\}, Pr[\omega] = p_{\omega}.$$

Probability Space: Formalism

Simplest physical model of a uniform probability space:



A bag of identical balls, except for their color (or a label). If the bag is well shaken, every ball is equally likely to be picked.

$$\Omega = \{ \text{white, red, yellow, grey, purple, blue, maroon, green} \} \\ Pr[\text{blue}] = \frac{1}{8}.$$

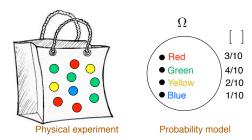
An important remark

- ightharpoonup The random experiment selects one and only one outcome in Ω .
- ► For instance, when we flip a fair coin twice

 - ▶ The experiment selects *one* of the elements of Ω .
- ▶ In this case, its wrong to think that $\Omega = \{H, T\}$ and that the experiment selects two outcomes.
- Why? Because this would not describe how the two coin flips are related to each other.
- ► For instance, say we glue the coins side-by-side so that they face up the same way. Then one gets *HH* or *TT* with probability 50% each. This is not captured by 'picking two outcomes.'

Probability Space: Formalism

Simplest physical model of a non-uniform probability space:



$$\Omega = \{ \text{Red, Green, Yellow, Blue} \}$$

 $Pr[\text{Red}] = \frac{3}{10}, Pr[\text{Green}] = \frac{4}{10}, \text{ etc.}$

Note: Probabilities are restricted to rational numbers: $\frac{N_k}{N}$.

Lecture 15: Summary

Modeling Uncertainty: Probability Space

- 1. Random Experiment
- 2. Probability Space: Ω ; $Pr[\omega] \in [0,1]$; $\sum_{\omega} Pr[\omega] = 1$.
- 3. Uniform Probability Space: $Pr[\omega] = 1/|\Omega|$ for all $\omega \in \Omega$.

CS70: On to Calculation.

Events, Conditional Probability, Independence, Bayes' Rule

- 1. Probability Basics Review
- 2. Events
- 3. Conditional Probability
- 4. Independence of Events
- 5. Bayes' Rule

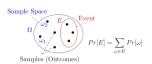
Probability of exactly one 'heads' in two coin flips?

Idea: Sum the probabilities of all the different outcomes that have exactly one 'heads': HT, TH.

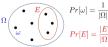
This leads to a definition!

Definition:

- ▶ An **event**, E, is a subset of outcomes: $E \subset \Omega$.
- ▶ The **probability of** *E* is defined as $Pr[E] = \sum_{\omega \in E} Pr[\omega]$.



Uniform Probability Space

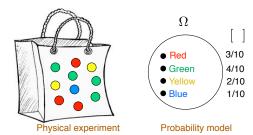


Probability Basics Review

Setup:

- Random Experiment. Flip a fair coin twice.
- Probability Space.
 - Sample Space: Set of outcomes, Ω . $\Omega = \{HH, HT, TH, TT\}$ (Note: Not $\Omega = \{H, T\}$ with two picks!)
 - ► **Probability:** $Pr[\omega]$ for all $\omega \in \Omega$. $Pr[HH] = \cdots = Pr[TT] = 1/4$
 - 1. $0 \le Pr[\omega] \le 1$.
 - 2. $\sum_{\omega \in \Omega} Pr[\omega] = 1$.

Event: Example



$$\Omega = \{ \text{Red, Green, Yellow, Blue} \}$$

$$Pr[\text{Red}] = \frac{3}{10}, Pr[\text{Green}] = \frac{4}{10}, \text{ etc.}$$

$$E = \{\textit{Red}, \textit{Green}\} \Rightarrow \textit{Pr}[E] = \frac{3+4}{10} = \frac{3}{10} + \frac{4}{10} = \textit{Pr}[\text{Red}] + \textit{Pr}[\text{Green}].$$

Set notation review





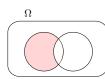
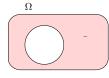
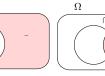


Figure : Two events

Figure : Union (or)

Figure : Difference (A, not B)





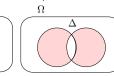


Figure : Complement (not)

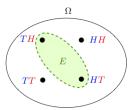
Figure : Intersection (and)

Figure : Symmetric difference (only one)

Probability of exactly one heads in two coin flips?

Sample Space, $\Omega = \{\textit{HH}, \textit{HT}, \textit{TH}, \textit{TT}\}.$

Uniform probability space: $Pr[HH] = Pr[HT] = Pr[TH] = Pr[TT] = \frac{1}{4}$. Event, E, "exactly one heads": $\{TH, HT\}$.



$$Pr[E] = \sum_{\omega \in E} Pr[\omega] = \frac{|E|}{|\Omega|} = \frac{2}{4} = \frac{1}{2}.$$

Example: 20 coin tosses.

20 coin tosses

Sample space: $\Omega = \text{set of 20 fair coin tosses}$.

$$\Omega = \{T, H\}^{20} \equiv \{0, 1\}^{20}; \ |\Omega| = 2^{20}.$$

What is more likely?

•
$$\omega_2 := (1,0,1,1,0,0,0,1,0,1,0,1,1,0,1,1,1,0,0,0)$$
?

Answer: Both are equally likely: $Pr[\omega_1] = Pr[\omega_2] = \frac{1}{|\Omega|}$.

▶ What is more likely?

 (E_1) Twenty Hs out of twenty, or

 (E_2) Ten Hs out of twenty?

Answer: Ten Hs out of twenty.

Why? There are many sequences of 20 tosses with ten Hs; only one with twenty Hs. $\Rightarrow Pr[E_1] = \frac{1}{|\Omega|} \ll Pr[E_2] = \frac{|E_2|}{|\Omega|}$.

$$|E_2| = {20 \choose 10} = 184,756.$$

Exactly 50 heads in 100 coin tosses.

Sample space: $\Omega=$ set of 100 coin tosses $=\{H,T\}^{100}$. $|\Omega|=2\times 2\times \cdots \times 2=2^{100}$.

Uniform probability space: $Pr[\omega] = \frac{1}{2^{100}}$.

Event E = "100 coin tosses with exactly 50 heads"

|E|

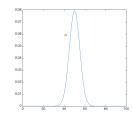
Choose 50 positions out of 100 to be heads.

 $|E| = \binom{100}{50}$.

$$Pr[E] = \frac{\binom{100}{50}}{2^{100}}.$$

Probability of *n* heads in 100 coin tosses.

$$\Omega = \{H, T\}^{100}; \ |\Omega| = 2^{100}.$$



Event E_n = 'n heads'; $|E_n| = \binom{100}{n}$

$$p_n := Pr[E_n] = \frac{|E_n|}{|\Omega|} = \frac{\binom{100}{n}}{2^{100}}$$

Observe:

- ► Concentration around mean: Law of Large Numbers;
- Bell-shape: Central Limit Theorem.

Calculation.

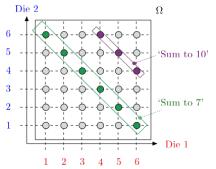
Stirling formula (for large *n*):

$$n! \approx \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$$
.

$$\binom{2n}{n} \approx \frac{\sqrt{4\pi n}(2n/e)^{2n}}{[\sqrt{2\pi n}(n/e)^n]^2} \approx \frac{4^n}{\sqrt{\pi n}}.$$

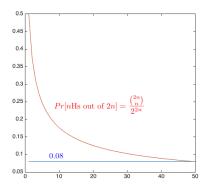
$$Pr[E] = \frac{|E|}{|\Omega|} = \frac{|E|}{2^{2n}} = \frac{1}{\sqrt{\pi n}} = \frac{1}{\sqrt{50\pi}} \approx .08.$$

Roll a red and a blue die.



$$Pr[\text{Sum to 7}] = \frac{6}{36}$$
 $Pr[\text{Sum to 10}] = \frac{3}{36}$

Exactly 50 heads in 100 coin tosses.



Probability is Additive

Theorem

(a) If events A and B are disjoint, i.e., $A \cap B = \emptyset$, then

$$Pr[A \cup B] = Pr[A] + Pr[B].$$

(b) If events $A_1, ..., A_n$ are pairwise disjoint, i.e., $A_k \cap A_m = \emptyset, \forall k \neq m$, then

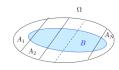
$$Pr[A_1 \cup \cdots \cup A_n] = Pr[A_1] + \cdots + Pr[A_n].$$

Proof:

Obvious.

Total probability

Assume that Ω is the union of the disjoint sets A_1, \ldots, A_N .



Then,

$$Pr[B] = Pr[A_1 \cap B] + \cdots + Pr[A_N \cap B].$$

Indeed, *B* is the union of the disjoint sets $A_n \cap B$ for n = 1, ..., N.

In "math": $\omega \in B$ is in exactly one of $A_i \cap B$.

Adding up probability of them, get $Pr[\omega]$ in sum.

..Did I say...

Add it up.

Consequences of Additivity

Theorem

(a) $Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$;

(inclusion-exclusion property)

(b)
$$Pr[A_1 \cup \cdots \cup A_n] \leq Pr[A_1] + \cdots + Pr[A_n];$$

(union bound)

(c) If $A_1, \dots A_N$ are a partition of Ω , i.e.,

pairwise disjoint and $\bigcup_{m=1}^{N} A_m = \Omega$, then

$$Pr[B] = Pr[B \cap A_1] + \cdots + Pr[B \cap A_N].$$

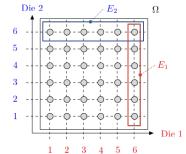
(law of total probability)

Proof:

(b) is obvious.

Proofs for (a) and (c)? Next...

Roll a Red and a Blue Die.



$$|E_1 \cup E_2| = |E_1| + |E_2| - |E_1 \cap E_2|$$

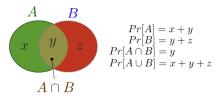
 E_1 = 'Red die shows 6'; E_2 = 'Blue die shows 6'

 $E_1 \cup E_2 =$ 'At least one die shows 6'

$$Pr[E_1] = \frac{6}{36}, Pr[E_2] = \frac{6}{36}, Pr[E_1 \cup E_2] = \frac{11}{36}.$$

Inclusion/Exclusion

$$Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$$



Another view. Any $\omega \in A \cup B$ is in $A \cap \overline{B}$, $A \cup B$, or $\overline{A} \cap B$. So, add it up.

Conditional probability: example.

Two coin flips. First flip is heads. Probability of two heads? $\Omega = \{HH, HT, TH, TT\}$; Uniform probability space. Event A = first flip is heads: $A = \{HH, HT\}$.



New sample space: A; uniform still.



Event B = two heads.

The probability of two heads if the first flip is heads.

The probability of B given A is 1/2.

A similar example.

Two coin flips. At least one of the flips is heads. \rightarrow Probability of two heads?

 $\Omega = \{\mathit{HH}, \mathit{HT}, \mathit{TH}, \mathit{TT}\}; \, \mathsf{uniform}.$ Event $A = \mathsf{at} \; \mathsf{least} \; \mathsf{one} \; \mathsf{flip} \; \mathsf{is} \; \mathsf{heads}. \; A = \{\mathit{HH}, \mathit{HT}, \mathit{TH}\}.$



New sample space: A; uniform still.

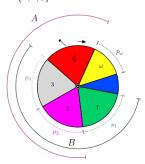


Event B = two heads.

The probability of two heads if at least one flip is heads. **The probability of** B **given** A is 1/3.

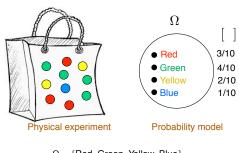
Yet another non-uniform example

Consider $\Omega = \{1, 2, ..., N\}$ with $Pr[n] = p_n$. Let $A = \{2, 3, 4\}, B = \{1, 2, 3\}$.



$$Pr[A|B] = \frac{p_2 + p_3}{p_1 + p_2 + p_3} = \frac{Pr[A \cap B]}{Pr[B]}$$

Conditional Probability: A non-uniform example

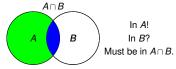


$$\begin{split} \Omega = & \{ \text{Red, Green, Yellow, Blue} \} \\ & \textit{Pr}[\text{Red} | \text{Red or Green}] = \frac{3}{7} = \frac{\textit{Pr}[\text{Red} \cap (\text{Red or Green})]}{\textit{Pr}[\text{Red or Green}]} \end{split}$$

Conditional Probability.

Definition: The **conditional probability** of B given A is

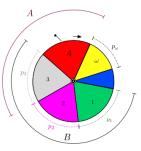
$$Pr[B|A] = \frac{Pr[A \cap B]}{Pr[A]}$$



$$Pr[B|A] = \frac{Pr[A \cap B]}{Pr[A]}.$$

Another non-uniform example

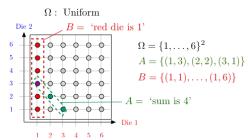
Consider $\Omega = \{1, 2, ..., N\}$ with $Pr[n] = p_n$. Let $A = \{3, 4\}, B = \{1, 2, 3\}$.



$$Pr[A|B] = \frac{p_3}{p_1 + p_2 + p_3} = \frac{Pr[A \cap B]}{Pr[B]}$$

More fun with conditional probability.

Toss a red and a blue die, sum is 4, What is probability that red is 1?



 $Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{3}$; versus Pr[B] = 1/6.

B is more likely given A.

Yet more fun with conditional probability.

Toss a red and a blue die, sum is 7, what is probability that red is 1?

$$Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{6}$$
; versus $Pr[B] = \frac{1}{6}$.

Observing A does not change your mind about the likelihood of B.

Product Rule

Recall the definition:

$$Pr[B|A] = \frac{Pr[A \cap B]}{Pr[A]}.$$

Hence,

$$Pr[A \cap B] = Pr[A]Pr[B|A].$$

Consequently,

$$Pr[A \cap B \cap C] = Pr[(A \cap B) \cap C]$$

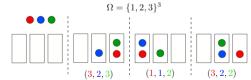
$$= Pr[A \cap B]Pr[C|A \cap B]$$

$$= Pr[A]Pr[B|A]Pr[C|A \cap B].$$

Emptiness..

Suppose I toss 3 balls into 3 bins.

A = ``1st bin empty''; B = ``2nd bin empty.'' What is Pr[A|B]?



 $\omega = (\text{bin of red ball}, \text{bin of blue ball}, \text{bin of green ball})$

$$Pr[B] = Pr[\{(a,b,c) \mid a,b,c \in \{1,3\}] = Pr[\{1,3\}^3] = \frac{8}{27}$$

$$Pr[A \cap B] = Pr[(3,3,3)] = \frac{1}{27}$$

$$Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]} = \frac{(1/27)}{(8/27)} = 1/8$$
; vs. $Pr[A] = \frac{8}{27}$.

 \emph{A} is less likely given \emph{B} : If second bin is empty the first is more likely to have balls in it.

Product Rule

Theorem Product Rule

Let $A_1, A_2, ..., A_n$ be events. Then

$$Pr[A_1 \cap \cdots \cap A_n] = Pr[A_1]Pr[A_2|A_1]\cdots Pr[A_n|A_1 \cap \cdots \cap A_{n-1}].$$

Proof: By induction.

Assume the result is true for n. (It holds for n = 2.) Then,

$$\begin{split} & Pr[A_1 \cap \dots \cap A_n \cap A_{n+1}] \\ & = Pr[A_1 \cap \dots \cap A_n] Pr[A_{n+1} | A_1 \cap \dots \cap A_n] \\ & = Pr[A_1] Pr[A_2 | A_1] \dots Pr[A_n | A_1 \cap \dots \cap A_{n-1}] Pr[A_{n+1} | A_1 \cap \dots \cap A_n], \end{split}$$

so that the result holds for n+1.

Gambler's fallacy.

Flip a fair coin 51 times.

A = "first 50 flips are heads"

B = "the 51st is heads"

Pr[B|A]?

 $\textit{A} = \{\textit{HH} \cdots \textit{HT}, \textit{HH} \cdots \textit{HH}\}$

 $B \cap A = \{HH \cdots HH\}$

Uniform probability space.

$$Pr[B|A] = \frac{|B \cap A|}{|A|} = \frac{1}{2}.$$

Same as Pr[B].

The likelihood of 51st heads does not depend on the previous flips.

Correlation

An example.

Random experiment: Pick a person at random.

Event A: the person has lung cancer.

Event B: the person is a heavy smoker.

Fact:

$$Pr[A|B] = 1.17 \times Pr[A].$$

Conclusion:

- ► Smoking increases the probability of lung cancer by 17%.
- ► Smoking causes lung cancer.

Correlation

Event *A*: the person has lung cancer. Event *B*: the person is a heavy smoker. $Pr[A|B] = 1.17 \times Pr[A]$.

A second look.

Note that

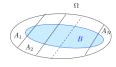
$$\begin{aligned} Pr[A|B] &= 1.17 \times Pr[A] &\Leftrightarrow & \frac{Pr[A \cap B]}{Pr[B]} = 1.17 \times Pr[A] \\ &\Leftrightarrow & Pr[A \cap B] = 1.17 \times Pr[A] Pr[B] \\ &\Leftrightarrow & Pr[B|A] = 1.17 \times Pr[B]. \end{aligned}$$

Conclusion:

- ▶ Lung cancer increases the probability of smoking by 17%.
- ► Lung cancer causes smoking. Really?

Total probability

Assume that Ω is the union of the disjoint sets A_1, \ldots, A_N .



Then.

$$Pr[B] = Pr[A_1 \cap B] + \cdots + Pr[A_N \cap B].$$

Indeed, *B* is the union of the disjoint sets $A_n \cap B$ for n = 1, ..., N. Thus,

$$Pr[B] = Pr[A_1]Pr[B|A_1] + \cdots + Pr[A_N]Pr[B|A_N].$$

Causality vs. Correlation

Events A and B are positively correlated if

$$Pr[A \cap B] > Pr[A]Pr[B]$$
.

(E.g., smoking and lung cancer.)

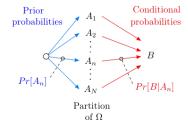
A and B being positively correlated does not mean that A causes B or that B causes A.

Other examples:

- ► Tesla owners are more likely to be rich. That does not mean that poor people should buy a Tesla to get rich.
- People who go to the opera are more likely to have a good career. That does not mean that going to the opera will improve your career.
- Rabbits eat more carrots and do not wear glasses. Are carrots good for eyesight?

Total probability

Assume that Ω is the union of the disjoint sets A_1, \ldots, A_N .



$$Pr[B] = Pr[A_1]Pr[B|A_1] + \cdots + Pr[A_N]Pr[B|A_N].$$

Proving Causality

Proving causality is generally difficult. One has to eliminate external causes of correlation and be able to test the cause/effect relationship (e.g., randomized clinical trials).

Some difficulties:

- A and B may be positively correlated because they have a common cause. (E.g., being a rabbit.)
- ▶ If *B* precedes *A*, then *B* is more likely to be the cause. (E.g., smoking.) However, they could have a common cause that induces *B* before *A*. (E.g., smart, CS70, Tesla.)

More about such questions later. For fun, check "N. Taleb: Fooled by randomness."

Is you coin loaded?

Your coin is fair w.p. 1/2 or such that Pr[H] = 0.6, otherwise.

You flip your coin and it yields heads.

What is the probability that it is fair?

Analysis:

A = 'coin is fair', B = 'outcome is heads'

We want to calculate P[A|B].

We know P[B|A] = 1/2, $P[B|\bar{A}] = 0.6$, $Pr[A] = 1/2 = Pr[\bar{A}]$ Now.

$$Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B] = Pr[A]Pr[B|A] + Pr[\bar{A}]Pr[B|\bar{A}]$$

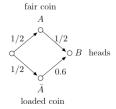
= (1/2)(1/2) + (1/2)0.6 = 0.55.

Thus,

$$Pr[A|B] = \frac{Pr[A]Pr[B|A]}{Pr[B]} = \frac{(1/2)(1/2)}{(1/2)(1/2) + (1/2)0.6} \approx 0.45.$$

Is you coin loaded?

A picture:



Imagine 100 situations, among which m := 100(1/2)(1/2) are such that A and B occur and n := 100(1/2)(0.6) are such that \bar{A} and B occur.

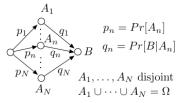
Thus, among the m+n situations where B occurred, there are mwhere A occurred.

Hence.

$$Pr[A|B] = \frac{m}{m+n} = \frac{(1/2)(1/2)}{(1/2)(1/2) + (1/2)0.6}.$$

Bayes Rule

Another picture: We imagine that there are N possible causes A_1,\ldots,A_N .



Imagine 100 situations, among which $100p_nq_n$ are such that A_n and B occur, for n = 1, ..., N.

Thus, among the $100 \sum_{m} p_{m} q_{m}$ situations where *B* occurred, there are $100p_nq_n$ where A_n occurred.

Hence.

$$Pr[A_n|B] = \frac{p_n q_n}{\sum_m p_m q_m}$$

Independence

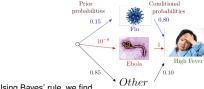
Definition: Two events A and B are **independent** if

$$Pr[A \cap B] = Pr[A]Pr[B]$$

Examples:

- ▶ When rolling two dice, A = sum is 7 and B = red die is 1 are
- ▶ When rolling two dice, A = sum is 3 and B = red die is 1 are not independent:
- ▶ When flipping coins, A = coin 1 yields heads and B = coin 2yields tails are independent:
- ▶ When throwing 3 balls into 3 bins, A = bin 1 is empty and B = bin 1bin 2 is empty are not independent:

Why do you have a fever?



Using Bayes' rule, we find

$$\textit{Pr}[\text{Flu}|\text{High Fever}] = \frac{0.15 \times 0.80}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 0.58$$

$$\textit{Pr}[\text{Ebola}|\text{High Fever}] = \frac{10^{-8} \times 1}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 5 \times 10^{-8}$$

$$\textit{Pr}[\text{Other}|\text{High Fever}] = \frac{0.85 \times 0.1}{0.15 \times 0.80 + 10^{-8} \times 1 + 0.85 \times 0.1} \approx 0.42$$

These are the posterior probabilities. One says that 'Flu' is the Most Likely a Posteriori (MAP) cause of the high fever.

Independence and conditional probability

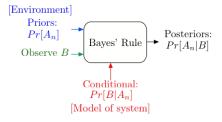
Fact: Two events A and B are independent if and only if

$$Pr[A|B] = Pr[A].$$

Indeed: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$, so that

$$Pr[A|B] = Pr[A] \Leftrightarrow \frac{Pr[A \cap B]}{Pr[B]} = Pr[A] \Leftrightarrow Pr[A \cap B] = Pr[A]Pr[B].$$

Bayes' Rule Operations



Bayes' Rule is the canonical example of how information changes our opinions.

Thomas Bayes



Source: Wikipedia.

Bayes Rule.



Using Bayes' rule, we find

$$P[A|B] = \frac{0.0016 \times 0.80}{0.0016 \times 0.80 + 0.9984 \times 0.10} = .013.$$

A 1.3% chance of prostate cancer with a positive PSA test.

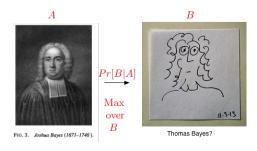
Surgery anyone?

Impotence...

Incontinence..

Death.

Thomas Bayes



A Bayesian picture of Thomas Bayes.

Summary

Events, Conditional Probability, Independence, Bayes' Rule

Key Ideas:

Conditional Probability:

$$Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$$

- ▶ Independence: $Pr[A \cap B] = Pr[A]Pr[B]$.
- ► Bayes' Rule:

$$Pr[A_n|B] = \frac{Pr[A_n]Pr[B|A_n]}{\sum_{m} Pr[A_m]Pr[B|A_m]}.$$

 $Pr[A_n|B] = posterior probability; Pr[A_n] = prior probability$.

▶ All these are possible:

$$Pr[A|B] < Pr[A]; Pr[A|B] > Pr[A]; Pr[A|B] = Pr[A].$$

Testing for disease.

Let's watch TV!!

Random Experiment: Pick a random male.

Outcomes: (test, disease)

A - prostate cancer.

B - positive PSA test.

- ightharpoonup Pr[A] = 0.0016, (.16 % of the male population is affected.)
- ightharpoonup Pr[B|A] = 0.80 (80% chance of positive test with disease.)
- ▶ $Pr[B|\overline{A}] = 0.10$ (10% chance of positive test without disease.)

From http://www.cpcn.org/01_psa_tests.htm and

http://seer.cancer.gov/statfacts/html/prost.html (10/12/2011.)

Positive PSA test (B). Do I have disease?

Pr[A|B]???