CS70: Jean Walrand: Lecture 30.

Linear Regression

#### CS70: Jean Walrand: Lecture 30.

#### Linear Regression

- 1. Preamble
- 2. Motivation for LR
- History of LR
- 4. Linear Regression
- 5. Derivation
- 6. More examples

The best guess about Y,

The best guess about Y, if we know only the distribution of Y, is

The best guess about Y, if we know only the distribution of Y, is E[Y].

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

The best guess about Y, if we know only the distribution of Y, is E[Y].

More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

#### **Proof:**

Let  $\hat{Y} := Y - E[Y]$ .

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ .

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

#### **Proof:**

Let  $\hat{Y} := Y - E[Y]$ . Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ .

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2]$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a$$
$$= E[\hat{Y}^2 + 2\hat{Y}c + c^2]$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a$$
$$= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a$$
$$= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2$$
$$= E[\hat{Y}^2] + 0 + c^2$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now, 
$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$
$$= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a$$
$$= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2$$
$$= E[\hat{Y}^2] + 0 + c^2 \ge E[\hat{Y}^2].$$

The best guess about Y, if we know only the distribution of Y, is E[Y]. More precisely, the value of a that minimizes  $E[(Y-a)^2]$  is a=E[Y].

Let 
$$\hat{Y} := Y - E[Y]$$
. Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now,

$$E[(Y-a)^{2}] = E[(Y-E[Y]+E[Y]-a)^{2}]$$

$$= E[(\hat{Y}+c)^{2}] \text{ with } c = E[Y]-a$$

$$= E[\hat{Y}^{2}+2\hat{Y}c+c^{2}] = E[\hat{Y}^{2}]+2E[\hat{Y}c]+c^{2}$$

$$= E[\hat{Y}^{2}]+0+c^{2} \ge E[\hat{Y}^{2}].$$

Hence, 
$$E[(Y - a)^2] \ge E[(Y - E[Y])^2], \forall a$$
.

Thus, if we want to guess the value of Y, we choose E[Y].

Thus, if we want to guess the value of Y, we choose E[Y]. Now assume we make some observation X related to Y.

Thus, if we want to guess the value of Y, we choose E[Y]. Now assume we make some observation X related to Y. How do we use that observation to improve our guess about Y?

Thus, if we want to guess the value of Y, we choose E[Y]. Now assume we make some observation X related to Y. How do we use that observation to improve our guess about Y? The idea is to use a function g(X) of the observation to estimate Y.

Thus, if we want to guess the value of Y, we choose E[Y].

Now assume we make some observation X related to Y.

How do we use that observation to improve our guess about Y?

The idea is to use a function g(X) of the observation to estimate Y.

The simplest function g(X) is a constant that does not depend of X.

Thus, if we want to guess the value of Y, we choose E[Y].

Now assume we make some observation X related to Y.

How do we use that observation to improve our guess about Y?

The idea is to use a function g(X) of the observation to estimate Y.

The simplest function g(X) is a constant that does not depend of X.

The next simplest function is linear: g(X) = a + bX.

Thus, if we want to guess the value of Y, we choose E[Y].

Now assume we make some observation X related to Y.

How do we use that observation to improve our guess about Y?

The idea is to use a function g(X) of the observation to estimate Y.

The simplest function g(X) is a constant that does not depend of X.

The next simplest function is linear: g(X) = a + bX.

What is the best linear function?

Thus, if we want to guess the value of Y, we choose E[Y].

Now assume we make some observation X related to Y.

How do we use that observation to improve our guess about Y?

The idea is to use a function g(X) of the observation to estimate Y.

The simplest function g(X) is a constant that does not depend of X.

The next simplest function is linear: g(X) = a + bX.

What is the best linear function? That is our next topic.

Thus, if we want to guess the value of Y, we choose E[Y].

Now assume we make some observation X related to Y.

How do we use that observation to improve our guess about Y?

The idea is to use a function g(X) of the observation to estimate Y.

The simplest function g(X) is a constant that does not depend of X.

The next simplest function is linear: g(X) = a + bX.

What is the best linear function? That is our next topic.

A bit later, we will consider a general function g(X).

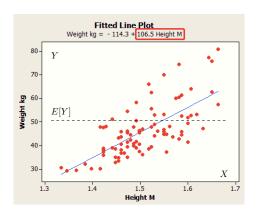
Example 1: 100 people.

Example 1: 100 people.

Let  $(X_n, Y_n)$  = (height, weight) of person n, for n = 1, ..., 100:

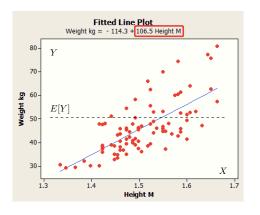
Example 1: 100 people.

Let  $(X_n, Y_n)$  = (height, weight) of person n, for n = 1, ..., 100:



Example 1: 100 people.

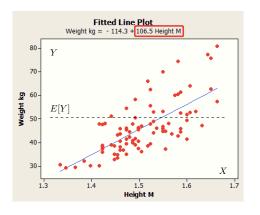
Let  $(X_n, Y_n)$  = (height, weight) of person n, for n = 1, ..., 100:



The blue line is Y = -114.3 + 106.5X. (X in meters, Y in kg.)

Example 1: 100 people.

Let  $(X_n, Y_n)$  = (height, weight) of person n, for n = 1, ..., 100:



The blue line is Y = -114.3 + 106.5X. (X in meters, Y in kg.) Best linear fit: Linear Regression.

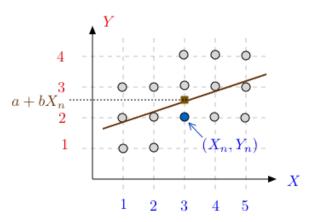
Example 2: 15 people.

Example 2: 15 people.

We look at two attributes:  $(X_n, Y_n)$  of person n, for n = 1, ..., 15:

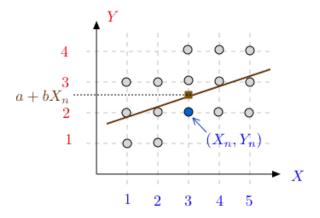
Example 2: 15 people.

We look at two attributes:  $(X_n, Y_n)$  of person n, for n = 1, ..., 15:



Example 2: 15 people.

We look at two attributes:  $(X_n, Y_n)$  of person n, for n = 1, ..., 15:



The line Y = a + bX is the linear regression.

**Definition** The covariance of *X* and *Y* is

$$cov(X, Y) := E[(X - E[X])(Y - E[Y])].$$

**Definition** The covariance of *X* and *Y* is

$$cov(X,Y) := E[(X - E[X])(Y - E[Y])].$$

$$cov(X, Y) = E[XY] - E[X]E[Y].$$

**Definition** The covariance of *X* and *Y* is

$$cov(X,Y) := E[(X - E[X])(Y - E[Y])].$$

**Fact** 

$$cov(X,Y) = E[XY] - E[X]E[Y].$$

$$E[(X - E[X])(Y - E[Y])] = E[XY - E[X]Y - XE[Y] + E[X]E[Y]]$$

**Definition** The covariance of *X* and *Y* is

$$cov(X,Y) := E[(X - E[X])(Y - E[Y])].$$

**Fact** 

$$cov(X, Y) = E[XY] - E[X]E[Y].$$

$$E[(X - E[X])(Y - E[Y])] = E[XY - E[X]Y - XE[Y] + E[X]E[Y]]$$
  
=  $E[XY] - E[X]E[Y] - E[X]E[Y] + E[X]E[Y]$ 

### **Definition** The covariance of *X* and *Y* is

$$cov(X,Y) := E[(X - E[X])(Y - E[Y])].$$

### **Fact**

$$cov(X,Y) = E[XY] - E[X]E[Y].$$

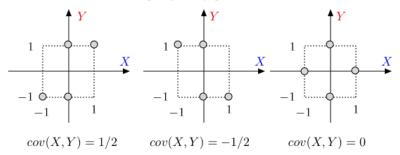
$$E[(X - E[X])(Y - E[Y])] = E[XY - E[X]Y - XE[Y] + E[X]E[Y]]$$

$$= E[XY] - E[X]E[Y] - E[X]E[Y] + E[X]E[Y]$$

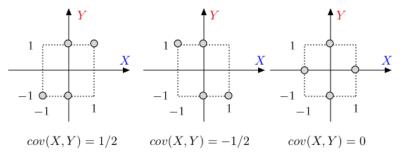
$$= E[XY] - E[X]E[Y].$$



### Four equally likely pairs of values

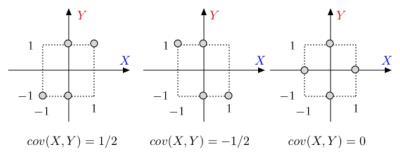


#### Four equally likely pairs of values



Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

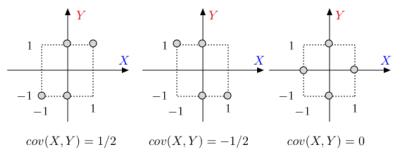
#### Four equally likely pairs of values



Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

When cov(X, Y) > 0, the RVs X and Y tend to be large or small together.

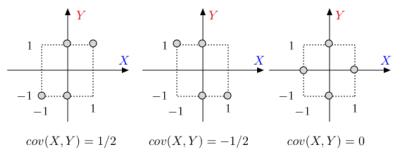
#### Four equally likely pairs of values



Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

When cov(X, Y) > 0, the RVs X and Y tend to be large or small together. X and Y are said to be positively correlated.

#### Four equally likely pairs of values

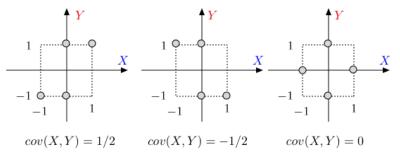


Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

When cov(X, Y) > 0, the RVs X and Y tend to be large or small together. X and Y are said to be positively correlated.

When cov(X, Y) < 0, when X is larger, Y tends to be smaller.

#### Four equally likely pairs of values

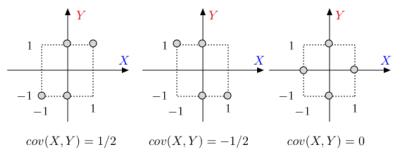


Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

When cov(X, Y) > 0, the RVs X and Y tend to be large or small together. X and Y are said to be positively correlated.

When cov(X, Y) < 0, when X is larger, Y tends to be smaller. X and Y are said to be negatively correlated.

#### Four equally likely pairs of values

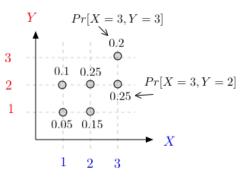


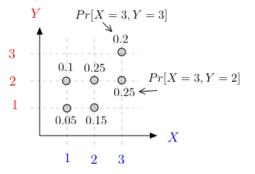
Note that E[X] = 0 and E[Y] = 0 in these examples. Then cov(X, Y) = E[XY].

When cov(X, Y) > 0, the RVs X and Y tend to be large or small together. X and Y are said to be positively correlated.

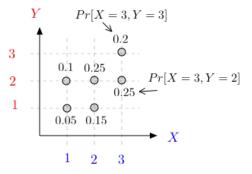
When cov(X, Y) < 0, when X is larger, Y tends to be smaller. X and Y are said to be negatively correlated.

When cov(X, Y) = 0, we say that X and Y are uncorrelated.

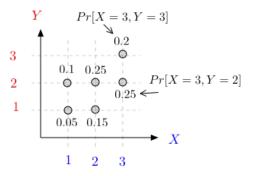




$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$



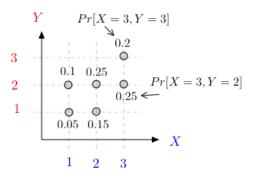
$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$
  
 $E[X^2] = 1^2 \times 0.15 + 2^2 \times 0.4 + 3^2 \times 0.45 = 5.8$ 



$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$

$$E[X^2] = 1^2 \times 0.15 + 2^2 \times 0.4 + 3^2 \times 0.45 = 5.8$$

$$E[Y] = 1 \times 0.2 + 2 \times 0.6 + 3 \times 0.2 = 2$$

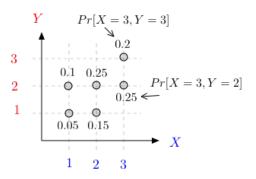


$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$

$$E[X^2] = 1^2 \times 0.15 + 2^2 \times 0.4 + 3^2 \times 0.45 = 5.8$$

$$E[Y] = 1 \times 0.2 + 2 \times 0.6 + 3 \times 0.2 = 2$$

$$E[XY] = 1 \times 0.05 + 1 \times 2 \times 0.1 + \dots + 3 \times 3 \times 0.2 = 4.85$$



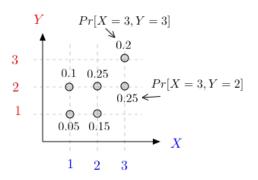
$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$

$$E[X^2] = 1^2 \times 0.15 + 2^2 \times 0.4 + 3^2 \times 0.45 = 5.8$$

$$E[Y] = 1 \times 0.2 + 2 \times 0.6 + 3 \times 0.2 = 2$$

$$E[XY] = 1 \times 0.05 + 1 \times 2 \times 0.1 + \dots + 3 \times 3 \times 0.2 = 4.85$$

$$cov(X, Y) = E[XY] - E[X]E[Y] = 1.05$$



$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$

$$E[X^{2}] = 1^{2} \times 0.15 + 2^{2} \times 0.4 + 3^{2} \times 0.45 = 5.8$$

$$E[Y] = 1 \times 0.2 + 2 \times 0.6 + 3 \times 0.2 = 2$$

$$E[XY] = 1 \times 0.05 + 1 \times 2 \times 0.1 + \dots + 3 \times 3 \times 0.2 = 4.85$$

$$cov(X, Y) = E[XY] - E[X]E[Y] = 1.05$$

$$var[X] = E[X^{2}] - E[X]^{2} = 2.19.$$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

### **Fact**

(a) var[X] = cov(X, X)

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) =$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

#### **Fact**

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

### **Proof:**

(a)-(b)-(c) are obvious.

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

#### **Fact**

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

- (a)-(b)-(c) are obvious.
- (d) In view of (c), one can subtract the means and assume that the RVs are zero-mean.

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

#### **Fact**

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

- (a)-(b)-(c) are obvious.
- (d) In view of (c), one can subtract the means and assume that the RVs are zero-mean. Then,

$$cov(aX + bY, cU + dV) = E[(aX + bY)(cU + dV)]$$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

#### **Fact**

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

- (a)-(b)-(c) are obvious.
- (d) In view of (c), one can subtract the means and assume that the RVs are zero-mean. Then,

$$cov(aX + bY, cU + dV) = E[(aX + bY)(cU + dV)]$$
$$= ac.E[XU] + ad.E[XV] + bc.E[YU] + bd.E[YV]$$

$$cov(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

#### **Fact**

- (a) var[X] = cov(X, X)
- (b) X, Y independent  $\Rightarrow cov(X, Y) = 0$
- (c) cov(a+X,b+Y) = cov(X,Y)
- (d) cov(aX + bY, cU + dV) = ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).

- (a)-(b)-(c) are obvious.
- (d) In view of (c), one can subtract the means and assume that the RVs are zero-mean. Then,

$$cov(aX + bY, cU + dV) = E[(aX + bY)(cU + dV)]$$

$$= ac.E[XU] + ad.E[XV] + bc.E[YU] + bd.E[YV]$$

$$= ac.cov(X, U) + ad.cov(X, V) + bc.cov(Y, U) + bd.cov(Y, V).$$

# Linear Regression: Non-Bayesian

### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, \dots, N\}$ ,

## Linear Regression: Non-Bayesian

### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2.$$

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2$$
.

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ .

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2$$
.

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ .

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2.$$

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ . The LR minimizes the sum of the squared errors.

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2.$$

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ . The LR minimizes the sum of the squared errors.

Why the squares and not the absolute values?

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2.$$

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ . The LR minimizes the sum of the squared errors.

Why the squares and not the absolute values? Main justification: much easier!

#### **Definition**

Given the samples  $\{(X_n, Y_n), n = 1, ..., N\}$ , the Linear Regression of Y over X is

$$\hat{Y} = a + bX$$

where (a, b) minimize

$$\sum_{n=1}^{N} (Y_n - a - bX_n)^2.$$

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ . The LR minimizes the sum of the squared errors.

Why the squares and not the absolute values? Main justification: much easier!

Note: This is a non-Bayesian formulation: there is no prior.

**Definition** 

# **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y],

#### **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

### **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

# **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X.

### **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a, b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X. The squared error is  $(Y - \hat{Y})^2$ .

#### **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X. The squared error is  $(Y - \hat{Y})^2$ . The LLSE minimizes the expected value of the squared error.

## **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X. The squared error is  $(Y - \hat{Y})^2$ . The LLSE minimizes the expected value of the squared error.

Why the squares and not the absolute values?

## **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X. The squared error is  $(Y - \hat{Y})^2$ . The LLSE minimizes the expected value of the squared error.

Why the squares and not the absolute values? Main justification: much easier!

## **Definition**

Given two RVs X and Y with known distribution Pr[X = x, Y = y], the Linear Least Squares Estimate of Y given X is

$$\hat{Y} = a + bX =: L[Y|X]$$

where (a,b) minimize

$$g(a,b) := E[(Y-a-bX)^2].$$

Thus,  $\hat{Y} = a + bX$  is our guess about Y given X. The squared error is  $(Y - \hat{Y})^2$ . The LLSE minimizes the expected value of the squared error.

Why the squares and not the absolute values? Main justification: much easier!

Note: This is a Bayesian formulation: there is a prior.

Observe that

$$\frac{1}{N}\sum_{n=1}^{N}(Y_n-a-bX_n)^2=E[(Y-a-bX)^2]$$

where one assumes that

$$(X, Y) = (X_n, Y_n), \text{ w.p. } \frac{1}{N} \text{ for } n = 1, ..., N.$$

Observe that

$$\frac{1}{N}\sum_{n=1}^{N}(Y_n-a-bX_n)^2=E[(Y-a-bX)^2]$$

where one assumes that

$$(X, Y) = (X_n, Y_n), \text{ w.p. } \frac{1}{N} \text{ for } n = 1, ..., N.$$

That is, the non-Bayesian LR is equivalent to the Bayesian LLSE that assumes that (X, Y) is uniform on the set of observed samples.

Observe that

$$\frac{1}{N}\sum_{n=1}^{N}(Y_n-a-bX_n)^2=E[(Y-a-bX)^2]$$

where one assumes that

$$(X, Y) = (X_n, Y_n), \text{ w.p. } \frac{1}{N} \text{ for } n = 1, ..., N.$$

That is, the non-Bayesian LR is equivalent to the Bayesian LLSE that assumes that (X, Y) is uniform on the set of observed samples.

Thus, we can study the two cases LR and LLSE in one shot.

Observe that

$$\frac{1}{N}\sum_{n=1}^{N}(Y_n-a-bX_n)^2=E[(Y-a-bX)^2]$$

where one assumes that

$$(X, Y) = (X_n, Y_n), \text{ w.p. } \frac{1}{N} \text{ for } n = 1, ..., N.$$

That is, the non-Bayesian LR is equivalent to the Bayesian LLSE that assumes that (X, Y) is uniform on the set of observed samples.

Thus, we can study the two cases LR and LLSE in one shot.

However, the interpretations are different!

# **Theorem**

# **Theorem**

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X]).$$

#### **Theorem**

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X]).$$
Proof 1:

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

Y - 
$$\hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also, 
$$E[(Y - \hat{Y})X] = 0$$
,

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y].$$
 Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X]).$ 
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra.

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

#### **Theorem**

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y-Y)(c+dX)]=0.$$

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .  
**Proof 1:**

Y - 
$$\hat{Y}$$
 =  $(Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$ . Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ . Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ ,

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,  
 $L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X, Y)}{var(X)}(X - E[X])$ .

# Proof 1:

Y - 
$$\hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c.d.

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

#### Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c.d. Now,

$$E[(Y - a - bX)^{2}] = E[(Y - \hat{Y} + \hat{Y} - a - bX)^{2}]$$

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

#### Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also.  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c, d. Now,

$$E[(Y-a-bX)^2] = E[(Y-\hat{Y}+\hat{Y}-a-bX)^2]$$
  
=  $E[(Y-\hat{Y})^2] + E[(\hat{Y}-a-bX)^2] + 0$ 

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

#### Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c, d. Now,

$$E[(Y-a-bX)^2] = E[(Y-\hat{Y}+\hat{Y}-a-bX)^2]$$
  
=  $E[(Y-\hat{Y})^2] + E[(\hat{Y}-a-bX)^2] + \frac{0}{2} \ge E[(Y-\hat{Y})^2].$ 

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

#### Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c, d. Now,

$$E[(Y-a-bX)^2] = E[(Y-\hat{Y}+\hat{Y}-a-bX)^2]$$
  
=  $E[(Y-\hat{Y})^2] + E[(\hat{Y}-a-bX)^2] + \frac{0}{2} \ge E[(Y-\hat{Y})^2].$ 

This shows that  $E[(Y - \hat{Y})^2] \le E[(Y - a - bX)^2]$ , for all (a,b).

#### Theorem

Consider two RVs X, Y with a given distribution

$$Pr[X = x, Y = y]$$
. Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

#### Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X])$$
. Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$$E[(Y - \hat{Y})(c + dX)] = 0$$
. Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some c, d. Now,

$$E[(Y-a-bX)^2] = E[(Y-\hat{Y}+\hat{Y}-a-bX)^2]$$
  
=  $E[(Y-\hat{Y})^2] + E[(\hat{Y}-a-bX)^2] + \frac{0}{2} \ge E[(Y-\hat{Y})^2].$ 

This shows that  $E[(Y - \hat{Y})^2] \le E[(Y - a - bX)^2]$ , for all (a, b). Thus  $\hat{Y}$  is the LLSE.

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ .

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because  $E[(Y - \hat{Y})E[X]] = 0$ .

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X,Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because  $E[(Y - \hat{Y})E[X]] = 0$ .

Now,

$$E[(Y - \hat{Y})(X - E[X])] = E[(Y - E[Y])(X - E[X])] - \frac{cov(X, Y)}{var[X]} E[(X - E[X])(X - E[X])]$$

$$Y - \hat{Y} = (Y - E[Y]) - \frac{cov(X, Y)}{var[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because  $E[(Y - \hat{Y})E[X]] = 0$ .

Now,

$$E[(Y - \hat{Y})(X - E[X])]$$

$$= E[(Y - E[Y])(X - E[X])] - \frac{cov(X, Y)}{var[X]} E[(X - E[X])(X - E[X])]$$

$$= (*) cov(X, Y) - \frac{cov(X, Y)}{var[X]} var[X] = 0. \quad \Box$$

(\*) Recall that 
$$cov(X, Y) = E[(X - E[X])(Y - E[Y])]$$
 and  $var[X] = E[(X - E[X])^2].$ 

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator?

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

$$E[|Y - L[Y|X]|^2] = E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^2]$$

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

$$E[|Y - L[Y|X]|^{2}] = E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^{2}]$$

$$= E[(Y - E[Y])^{2}] - 2(cov(X, Y)/var(X))E[(Y - E[Y])(X - E[X])]$$

$$+(cov(X, Y)/var(X))^{2}E[(X - E[X])^{2}]$$

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

$$\begin{split} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(cov(X, Y)/var(X))E[(Y - E[Y])(X - E[X])] \\ &+ (cov(X, Y)/var(X))^2 E[(X - E[X])^2] \\ &= var(Y) - \frac{cov(X, Y)^2}{var(X)}. \end{split}$$

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

We find

$$\begin{split} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(cov(X, Y)/var(X))E[(Y - E[Y])(X - E[X])] \\ &+ (cov(X, Y)/var(X))^2 E[(X - E[X])^2] \\ &= var(Y) - \frac{cov(X, Y)^2}{var(X)}. \end{split}$$

Without observations, the estimate is E[Y] = 0.

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

We find

$$\begin{split} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(cov(X, Y)/var(X))E[(Y - E[Y])(X - E[X])] \\ &+ (cov(X, Y)/var(X))^2 E[(X - E[X])^2] \\ &= var(Y) - \frac{cov(X, Y)^2}{var(X)}. \end{split}$$

Without observations, the estimate is E[Y] = 0. The error is var(Y).

We saw that the LLSE of Y given X is

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X]).$$

How good is this estimator? That is, what is the mean squared estimation error?

We find

$$\begin{split} E[|Y - L[Y|X]|^2] &= E[(Y - E[Y] - (cov(X, Y)/var(X))(X - E[X]))^2] \\ &= E[(Y - E[Y])^2] - 2(cov(X, Y)/var(X))E[(Y - E[Y])(X - E[X])] \\ &+ (cov(X, Y)/var(X))^2 E[(X - E[X])^2] \\ &= var(Y) - \frac{cov(X, Y)^2}{var(X)}. \end{split}$$

Without observations, the estimate is E[Y] = 0. The error is var(Y). Observing X reduces the error.

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X])$$

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = var(Y) - \frac{cov(X, Y)^2}{var(X)}.$$

We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X])$$

and

$$E[|Y - L[Y|X]|^2] = var(Y) - \frac{cov(X, Y)^2}{var(X)}.$$

Here is a picture when E[X] = 0, E[Y] = 0:

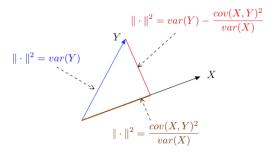
We saw that

$$L[Y|X] = \hat{Y} = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X])$$

and

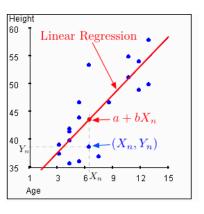
$$E[|Y - L[Y|X]|^2] = var(Y) - \frac{cov(X, Y)^2}{var(X)}.$$

Here is a picture when E[X] = 0, E[Y] = 0:



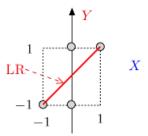
Example 1:

#### Example 1:

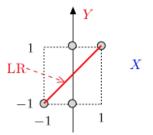


Example 2:

Example 2:

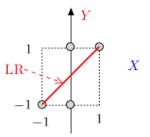


#### Example 2:



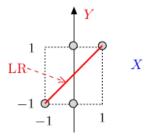
$$E[X] =$$

Example 2:



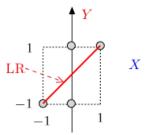
$$E[X] = 0;$$

#### Example 2:



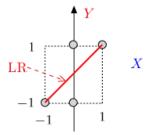
$$E[X] = 0; E[Y] =$$

#### Example 2:



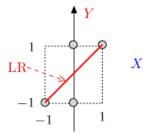
$$E[X] = 0; E[Y] = 0;$$

#### Example 2:



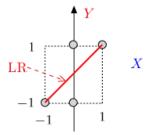
$$E[X] = 0; E[Y] = 0; E[X^2] =$$

#### Example 2:



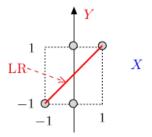
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2;$$

#### Example 2:



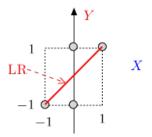
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] =$$

#### Example 2:



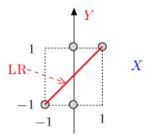
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$

#### Example 2:



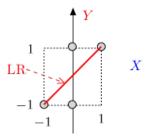
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 =$ 

#### Example 2:



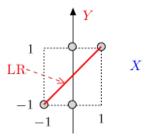
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2;$ 

#### Example 2:



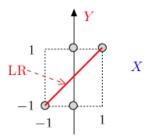
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] =$ 

#### Example 2:



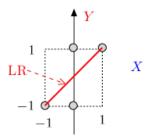
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] = 1/2;$ 

#### Example 2:



$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] = 1/2;$   
 $LR: \hat{Y} = E[Y] + \frac{cov(X, Y)}{var[X]}(X - E[X]) =$ 

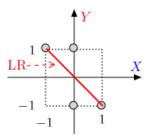
#### Example 2:



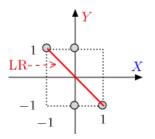
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = 1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] = 1/2;$   
 $LR: \hat{Y} = E[Y] + \frac{cov(X, Y)}{var[X]}(X - E[X]) = X.$ 

Example 3:

#### Example 3:

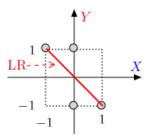


#### Example 3:



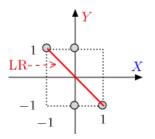
$$E[X] =$$

#### Example 3:



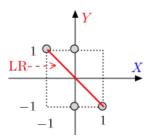
$$E[X] = 0;$$

#### Example 3:



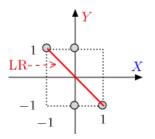
$$E[X] = 0; E[Y] =$$

#### Example 3:



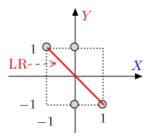
$$E[X] = 0; E[Y] = 0;$$

#### Example 3:



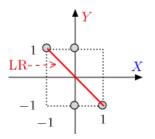
$$E[X] = 0; E[Y] = 0; E[X^2] =$$

#### Example 3:



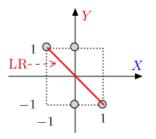
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2;$$

#### Example 3:



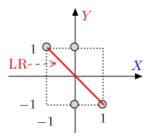
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] =$$

#### Example 3:



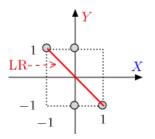
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$

#### Example 3:



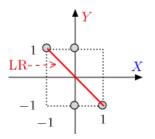
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 =$ 

#### Example 3:



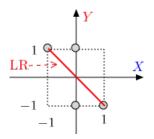
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2;$ 

#### Example 3:



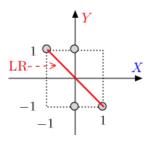
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] =$ 

#### Example 3:



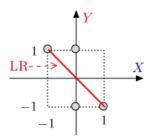
$$E[X] = 0; E[Y] = 0; E[X^2] = 1/2; E[XY] = -1/2;$$
  
 $var[X] = E[X^2] - E[X]^2 = 1/2; cov(X, Y) = E[XY] - E[X]E[Y] = -1/2;$ 

#### Example 3:



$$E[X] = 0$$
;  $E[Y] = 0$ ;  $E[X^2] = 1/2$ ;  $E[XY] = -1/2$ ;  $var[X] = E[X^2] - E[X]^2 = 1/2$ ;  $cov(X, Y) = E[XY] - E[X]E[Y] = -1/2$ ;  $LR: \hat{Y} = E[Y] + \frac{cov(X, Y)}{var[X]}(X - E[X]) =$ 

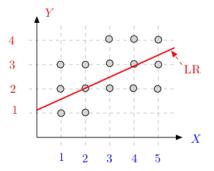
#### Example 3:



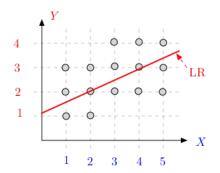
$$E[X] = 0$$
;  $E[Y] = 0$ ;  $E[X^2] = 1/2$ ;  $E[XY] = -1/2$ ;  $var[X] = E[X^2] - E[X]^2 = 1/2$ ;  $cov(X, Y) = E[XY] - E[X]E[Y] = -1/2$ ;  $LR: \hat{Y} = E[Y] + \frac{cov(X, Y)}{var[X]}(X - E[X]) = -X$ .

Example 4:

Example 4:

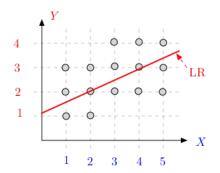


Example 4:



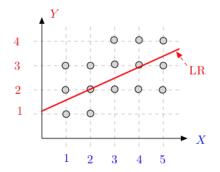
$$E[X] =$$

Example 4:



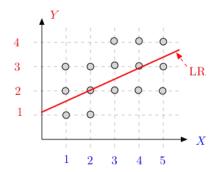
$$E[X] = 3;$$

Example 4:



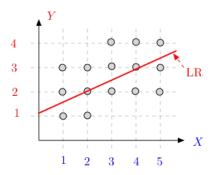
$$E[X] = 3; E[Y] =$$

Example 4:



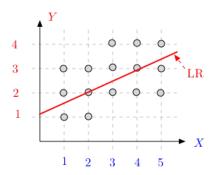
$$E[X] = 3; E[Y] = 2.5;$$

Example 4:



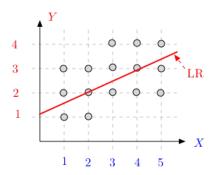
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1+2^2+3^2+4^2+5^2) = 11;$$

Example 4:



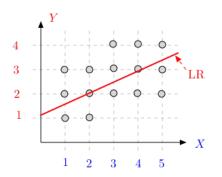
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$
  
 $E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$ 

Example 4:



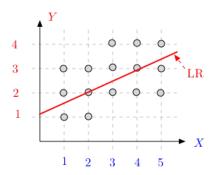
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$
  
 $E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$   
 $var[X] = 11 - 9 = 2;$ 

Example 4:



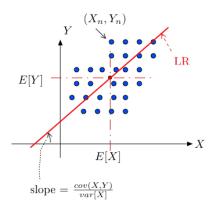
$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$
  
 $E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$   
 $var[X] = 11 - 9 = 2; cov(X, Y) = 8.4 - 3 \times 2.5 = 0.9;$ 

Example 4:

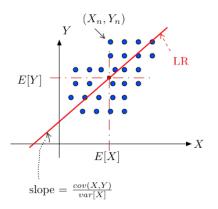


$$E[X] = 3; E[Y] = 2.5; E[X^2] = (3/15)(1 + 2^2 + 3^2 + 4^2 + 5^2) = 11;$$
  
 $E[XY] = (1/15)(1 \times 1 + 1 \times 2 + \dots + 5 \times 4) = 8.4;$   
 $var[X] = 11 - 9 = 2; cov(X, Y) = 8.4 - 3 \times 2.5 = 0.9;$   
 $LR: \hat{Y} = 2.5 + \frac{0.9}{2}(X - 3) = 1.15 + 0.45X.$ 

### LR: Another Figure



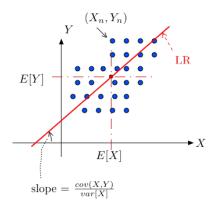
### LR: Another Figure



#### Note that

▶ the LR line goes through (E[X], E[Y])

# LR: Another Figure



#### Note that

- ▶ the LR line goes through (E[X], E[Y])
- ▶ its slope is  $\frac{cov(X,Y)}{var(X)}$

Linear Regression

**Linear Regression** 

1. Linear Regression:  $L[Y|X] = E[Y] + \frac{cov(X,Y)}{var(X)}(X - E[X])$ 

**Linear Regression** 

- 1. Linear Regression:  $L[Y|X] = E[Y] + \frac{cov(X,Y)}{var(X)}(X E[X])$
- 2. Non-Bayesian: minimize  $\sum_n (Y_n a bX_n)^2$

#### **Linear Regression**

- 1. Linear Regression:  $L[Y|X] = E[Y] + \frac{cov(X,Y)}{var(X)}(X E[X])$
- 2. Non-Bayesian: minimize  $\sum_{n} (Y_n a bX_n)^2$
- 3. Bayesian: minimize  $E[(Y-a-bX)^2]$